

A. Stochastic Processes

A (Discrete Time) stochastic process is a simple sequence  $X_0, X_1, X_2, \dots$  of random variables defined on some fixed probability triple  $(\Omega, \mathcal{F}, P)$ . The random variables  $\{X_n\}$  are typically not independent. In this context we often think of  $n$  as representing time, and  $X_n$  represents the value of a random quantity at the time  $n$ .

Theorem(A1): Let  $U$  be a random variable whose distribution is Lebesgue measure on  $[0, 1]$ . Let  $F$  be any cumulative distribution function, and set  $\phi(u) = \inf \{x; F(x) \geq u\}$  for  $0 < u < 1$ . Then  $P(\phi(U) \leq x) = F(x)$  for each  $x \in \mathbb{R}$ ; in words, the cumulative distribution function of  $\phi(U)$  is  $F$ .

Proof: Since  $F$  is right-continuous (cumulative) we can say that  $\inf \{x; F(x) \geq u\} = \min \{x; F(x) \geq u\}$ . It will follow that  $\phi(u) \leq x \iff u \leq F(x)$ . Hence, since  $0 \leq F(x) \leq 1$ , we obtain that  $P(\phi(U) \leq x) = P(U \leq F(x)) = F(x)$ .  $\square$

Theorem(A2): Let  $\mu_1, \mu_2, \mu_3, \dots$  be any sequence of Borel probability measures on  $\mathbb{R}$ . Then there exists a probability space  $(\Omega, \mathcal{F}, P)$  and random variables  $X_1, X_2, X_3, \dots$  defined on  $(\Omega, \mathcal{F}, P)$  such  $\{X_n\}$  are independent and  $L(X_n) = \mu_n$ .

Proof: Let  $(\Omega, \mathcal{F}, P)$  be infinite ~~at~~ independent fair coin tossing so that  $r_1, r_2, r_3, \dots$  are i.i.d. with  $P(r_i = 0) = P(r_i = 1) = \frac{1}{2}$ . Let  $\{Z_{ij}\}$  be a two-dimensional array filled by these  $r_i$  as follows:

$$\begin{pmatrix} Z_{11} & Z_{12} & Z_{13} & \dots \\ Z_{21} & Z_{22} & Z_{23} & \dots \\ Z_{31} & Z_{32} & Z_{33} & \dots \\ \vdots & \vdots & \vdots & \dots \end{pmatrix} \equiv \begin{pmatrix} r_1 & r_3 & r_6 & \dots \\ r_2 & r_5 & \dots & \dots \\ r_4 & \dots & \dots & \dots \\ \vdots & \vdots & \vdots & \dots \end{pmatrix}$$

Hence  $\{Z_{ij}\}$  are independent with  $P(Z_{ij} = 0) = P(Z_{ij} = 1) = \frac{1}{2}$ .

Then for each  $n \in \mathbb{N}$  we set  $U_n = \sum_{k=1}^{\infty} \frac{Z_{nk}}{2^k}$ . Since  $Z_{ij}$  are independent,  $U_n$  is also independent. Also by the way the  $U_n$  were constructed we have

$$P\left(\frac{j}{2^k} \leq U_n < \frac{j+1}{2^k}\right) = \frac{1}{2^k}, \quad k \in \mathbb{N} \text{ and } 0 \leq j < 2^k.$$

By the additivity and continuity of probabilities, this implies that  $P(a \leq U_n < b) = b - a$ , whenever  $0 \leq a < b \leq 1$ . Therefore, each  $U_n$  follows a uniform distribution (Lebesgue Measure) on  $[0, 1]$ .

Finally, we set  $F_n(x) = \gamma_n(-\infty, x]$  for  $x \in \mathbb{R}$  and set  $\Phi_n(u) = \inf \{x \mid u \leq F_n(x)\}$  for  $0 < u < 1$  and set  $X_n = \Phi_n(U_n)$ . Then  $\{X_n\}$  are independent by definition since  $U_n$  are independent. And by Theorem (1) we get  $X_n \sim \gamma_n$ , as required.

Theorem (A3): (The Bounded Convergence Theorem) Let  $\{X_n\}$  be a sequence of random variables, with  $\lim X_n = X$ . Suppose there is  $K \in \mathbb{R}$  such that  $|X_n| \leq K$  for all  $n \in \mathbb{N}$ . Then  $E(X) = \lim E(X_n)$ .

Proof: We have from triangle inequality that,

$$|E(X) - E(X_n)| = |E(X - X_n)| \leq E(|X - X_n|).$$

We will show that  $E(|X - X_n|) \rightarrow 0$  as  $n \rightarrow \infty$ . Thus for  $\epsilon > 0$  we set,

$$A_n = \{ \omega \in \Omega ; |X(\omega) - X_n(\omega)| > \epsilon \}.$$
 Then we have,

$$|X(\omega) - X_n(\omega)| \leq \epsilon + 2K \mathbb{1}_{A_n}(\omega), \text{ so that}$$

$$E(|X - X_n|) \leq \epsilon + 2K P(A_n).$$

However we have the following,

$$\begin{aligned} \limsup_n E(|X - X_n|) &\leq \epsilon + 2K \limsup P(A_n) \\ &\leq \epsilon + 2K P(\limsup A_n) \\ &= \epsilon, \end{aligned}$$

Since  $|X(\omega) - X_n(\omega)| \rightarrow 0$  for all  $\omega \in \Omega$  as  $n \rightarrow \infty$ .

Therefore  $\limsup_{n \rightarrow \infty} A_n$  is the empty set. Hence we get

$E(|X - X_n|) \rightarrow 0$  as claimed. □

## B. Markov Chains

A Markov Chain consists of three ingredients: a state space  $S$  which is a finite or a countable set, an initial distribution  $\{v_i\}_{i \in S}$  and transition probabilities  $\{P_{ij}\}_{i,j \in S}$ ,  $\sum_j P_{ij} = 1$ . We can say that

$$P(X_1=j) = \sum_{i \in S} P(X_0=i, X_1=j) = \sum_{i \in S} v_i P_{ij}.$$

Theorem (B1): (Markov Chain Existence Theorem) Given a non-empty countable set  $S$ , and non-negative numbers  $\{v_i\}_{i \in S}$  and  $\{P_{ij}\}_{i,j \in S}$ , with  $\sum_j v_j = 1$  and  $\sum_j P_{ij} = 1$  for each  $i \in S$ , there exists a probability triple  $(\Omega, \mathcal{F}, P)$  and random variables  $X_0, X_1, \dots$  defined on  $(\Omega, \mathcal{F}, P)$  such that

$$P(X_0=i_0, X_1=i_1, \dots, X_n=i_n) = v_{i_0} P_{i_0 i_1} \dots P_{i_{n-1} i_n}.$$

for all  $n \in \mathbb{N}$  and all  $i_0, \dots, i_n \in S$ .

Proof: We let  $(\Omega, \mathcal{F}, P)$  be a Lebesgue measure on  $[0, 1]$ .

We construct the random variables  $\{X_n\}$  as follows:

1. Partition  $[0, 1]$  into intervals  $\{I_i^{(0)}\}_{i \in S}$  with  $\text{length}(I_i^{(0)}) = v_i$
2. Partition each  $I_i^{(0)}$  into intervals  $\{I_{ij}^{(1)}\}_{j \in S}$  with  $\text{length}(I_{ij}^{(1)}) = v_i P_{ij}$ .
3. Inductively partition  $[0, 1]$  into intervals  $\{I_{i_0 i_1 \dots i_n}^{(n)}\}_{i_0, \dots, i_n \in S}$  such that  $I_{i_0, \dots, i_n}^{(n)} \subseteq I_{i_0, \dots, i_{n-1}}^{(n-1)}$  for all  $n \in \mathbb{N}$
4. Define  $X_n$  saying that  $X_n(\omega) = i_n$  if  $\omega \in I_{i_0 i_1 \dots i_n}^{(n)}$  for some choice of  $i_0, \dots, i_{n-1} \in S \implies \{X_n\}$  have desired properties  $\square$