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D. Weak Convergence

Def. Given Borel probability distributions $\mu, \mu_1, \mu_2, \mu_3, \dots$ on \mathbb{R} we shall write $\mu_n \rightarrow \mu$ and say that $\{\mu_n\}$ converges weakly to μ , if

$$\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu \text{ for all bounded continuous functions } f: \mathbb{R} \rightarrow \mathbb{R}.$$

Theorem 1: The following statements are equivalent;

- (1) $\mu_n \rightarrow \mu$ ($\{\mu_n\}$ converges weakly to μ)
- (2) $\mu_n(A) \rightarrow \mu(A)$ for all measurable sets A such that $\mu(\partial A) = 0$
- (3) $\mu_n((-\infty, x]) \rightarrow \mu((-\infty, x])$ for all $x \in \mathbb{R}$ such that $\mu\{x\} = 0$
- (4) (Skorohod's Theorem) there are random variables Y, Y_1, Y_2, \dots defined jointly on some probability triple, with $L(Y) = \mu$ and $L(Y_n) = \mu_n$ for each $n \in \mathbb{N}$ such that $Y_n \rightarrow Y$ with probability 1.
- (5) $\int_{\mathbb{R}} f d\mu_n \rightarrow \int_{\mathbb{R}} f d\mu$ for all bounded Borel-measurable functions $f: \mathbb{R} \rightarrow \mathbb{R}$ such that $\mu(D_f) = 0$, where D_f is the set of points where f is discontinuous.

Proof:

(5) \rightarrow (1): follows quickly from the definition of weak convergence.

(5) \Rightarrow (2): This follows by setting $f = 1_A$ so that we get

$Df = \partial A$. Thereby $\gamma(Pf) = \gamma(\partial A) = 0$. Hence we get that

$$\gamma_n(A) = \int_{\mathbb{R}} f d\gamma_n \rightarrow \int_{\mathbb{R}} f d\gamma = \gamma(A).$$

(1) \Rightarrow (3): Let $\epsilon > 0$ and let f be the function defined by

$f(t) = 1$ for $t \leq x$ and $f(t) = 0$ for $t \geq x + \epsilon$. And f is linear on the interval $(x, x + \epsilon)$. Therefore we have $1_{(-\infty, x]} \leq f \leq 1_{(-\infty, x + \epsilon]}$.

Hence,

$$\limsup_{n \rightarrow \infty} \gamma_n((-\infty, x]) \leq \limsup_{n \rightarrow \infty} \int f d\gamma_n = \int f d\gamma \leq \gamma((-\infty, x + \epsilon]).$$

Since this is true for any $\epsilon > 0$ we conclude that

$$\limsup_{n \rightarrow \infty} \gamma_n((-\infty, x]) \leq \gamma((-\infty, x]).$$

Similarly, if we let f be the function defined by $f(t) = 1$ for $t \leq x - \epsilon$ and $f(t) = 0$, for $t \geq x$, with f linear on the interval

$(x - \epsilon, x)$. Then we will get $1_{(-\infty, x - \epsilon]} \leq f \leq 1_{(-\infty, x]}$ and we get,

$$\liminf_{n \rightarrow \infty} \gamma_n((-\infty, x_n]) \geq \liminf_{n \rightarrow \infty} \int f d\gamma_n = \int f d\gamma \geq \gamma((-\infty, x - \epsilon]).$$

Since this is true for all $\epsilon > 0$, we get

$$\liminf_{n \rightarrow \infty} \gamma_n((-\infty, x_n]) \geq \gamma((-\infty, x)).$$

Since $\gamma\{x\} = 0 \Rightarrow \gamma((-\infty, x]) = \gamma((-\infty, x))$. We get as desired,

$$\limsup_{n \rightarrow \infty} \gamma_n((-\infty, x_n]) = \liminf_{n \rightarrow \infty} \gamma_n((-\infty, x_n]) = \gamma((-\infty, x]).$$

(3) \Rightarrow (4):

We first define the cumulative distribution functions, by

$F_n(x) = \mu_n((-\infty, x])$ and $F(x) = \mu((-\infty, x])$. Then we let

(Ω, \mathcal{F}, P) be a Lebesgue measure on $[0, 1]$ and let

$Y_n(\omega) = \inf \{x; F_n(x) \geq \omega\}$ and $Y(\omega) = \inf \{x; F(x) \geq \omega\}$,

then as in Theorem (A1) we have $L(Y_n) = \mu_n$ and $L(Y) = \mu$.

So, $F(z) < a \implies z \leq Y(a)$

$F(z) \geq b \implies z \geq Y(b)$

Since $\{F_n\} \rightarrow F$ at most points it seems that $\{Y_n\} \rightarrow Y$ at most points. We will prove that $\{Y_n\} \rightarrow Y$ at the points of continuity of Y . Since Y is non-decreasing we can have at most a countable number of discontinuities: indeed it will at most $m(Y(n+1) - Y(n)) < \infty$ discontinuities of size $\geq 1/m$ within the interval $(n, n+1]$. However since countable sets have measure 0 this implies that $\{Y_n\} \rightarrow Y$ with probability 1.

Suppose then that F is continuous at w and let $y = Y(w)$. For any $\epsilon > 0$ we claim that $F(y - \epsilon) < w < F(y + \epsilon)$. If we had $F(y - \epsilon) = w$ then setting $w = y - \epsilon$ and $b = w$, this would imply that $Y(w) \leq y - \epsilon = Y(w) - \epsilon$, a contradiction. Or if we had $F(y + \epsilon) = w$ then setting $z = y + \epsilon$ and $a = w + \delta$ would imply that $Y(w + \delta) \geq y + \epsilon = Y(w) + \epsilon$ for all $\delta > 0$ contradicting the continuity of Y at w . So we must have $F(y - \epsilon) < w < F(y + \epsilon)$, $\forall \epsilon > 0$.

Now, given that $\epsilon > 0$ we find $\epsilon' > 0$ $0 < \epsilon' < \epsilon$ such that $\chi_{[y-\epsilon', y+\epsilon']}$ $\rightarrow 0$. Then $F_n(y-\epsilon') \rightarrow F(y-\epsilon')$ and $F_n(y+\epsilon') \rightarrow F(y+\epsilon')$, so $F_n(y-\epsilon') < w < F_n(y+\epsilon')$ for all sufficiently large n . This will imply that $y-\epsilon' \leq Y_n(w) \leq y+\epsilon'$ i.e. $|Y_n(w) - Y(w)| \leq \epsilon' < \epsilon$ for all sufficiently large n .

Hence $Y_n(w) \rightarrow Y(w)$.

(4) \Rightarrow (5): We know that if f is continuous at x and we have $\{x_n\} \rightarrow x$, then $f(x_n) \rightarrow f(x)$. Hence if $\{Y_n\} \rightarrow Y$ and $Y \notin D_f$ then $\{f(Y_n)\} \rightarrow f(Y)$.

Then we can say that $P[f(Y_n) \rightarrow f(Y)] \geq P[Y_n \rightarrow Y, Y \notin D_f]$.

But by assumption we have $P(Y_n \rightarrow Y) = 1$ and $P(Y \notin D_f) =$

$P(Y \in D_f^c) = \nu(P_f^c) = 1$. So $P(f(Y_n) \rightarrow f(Y)) = 1$. If f

is bounded then by Bounded Convergence Theorem, we get

~~$E(f) = E[f(Y_n)] \rightarrow E[f(Y)]$~~ i.e. $\int f d\mu_n \rightarrow \int f d\mu$. \square

Example. We suppose that X_1, X_2, X_3, \dots are i.i.d random variables with a finite mean m , and $S_n = \frac{1}{n}(X_1 + X_2 + X_3 + \dots + X_n)$. Then the Weak Law of Large Numbers says that $P(S_n \leq m - \epsilon) \rightarrow 0$ and $P(S_n \leq m + \epsilon) \rightarrow 1$ as $n \rightarrow \infty$. This states that $L(S_n) \Rightarrow \delta_m(\cdot)$, a point mass at m .

Theorem 2: If $\{X_n\} \rightarrow X$ in probability, then $L(X_n) \Rightarrow L(X)$.

Proof: For any $\epsilon > 0$, if we have $X > z + \epsilon$ and $|X_n - X| < \epsilon$, then

we must have $X_n > z$. That is, $\{X > z + \epsilon\} \cap \{|X_n - X| < \epsilon\}$

$\subseteq \{X_n > z\}$. Then taking complements of the statement we

get $\{X \leq z + \epsilon\} \cup \{|X_n - X| > \epsilon\} \supseteq \{X_n \leq z\}$. Hence we get,

$P[X_n \leq z] \leq P[X \leq z + \epsilon] + P[|X_n - X| \geq \epsilon]$.

Since $\{X_n\} \rightarrow X$ in probability we get that,

$$\limsup_{n \rightarrow \infty} P(X_n \leq z) \leq P(X \leq z + \epsilon). \text{ Letting } \epsilon \rightarrow 0 \text{ gives}$$

$$\limsup_{n \rightarrow \infty} P(X_n \leq z) \leq P(X \leq z).$$

Similarly interchanging X and X_n and replacing z with $z - \epsilon$, in

the above gives $P(X \leq z - \epsilon) \leq P(X_n \leq z) + P(|X_n - X| \geq \epsilon)$, or

$P(X_n \leq z) \geq P(X \leq z - \epsilon) - P(|X_n - X| \geq \epsilon)$, so we get

$\liminf P(X_n \leq z) \geq P(X \leq z - \epsilon)$. Letting $\epsilon \rightarrow 0$ gives us

$$\liminf_{n \rightarrow \infty} P(X_n \leq z) \geq P(X \leq z).$$

If $P(X = z) = 0$, we must have $P(X \leq z) = P(X < z)$. So

should get $\limsup_{n \rightarrow \infty} P(X_n \leq z) = \liminf_{n \rightarrow \infty} P(X_n \leq z) = P(X \leq z)$. \square

Theorem B3: Suppose that $Z(X_n) \Rightarrow Z(X)$ with $X_n \geq 0$. Then

$$E(X) \leq \liminf E(X_n).$$

Proof: By Skorohod's Theorem, we can find random variables

Y_n and Y with $Z(Y_n) = Z(X_n)$, $Z(Y) = Z(X)$ and $Y_n \rightarrow Y$

with probability 1. Then by Fatou's Lemma,

$$E(X) = E(Y) = E(\liminf Y_n) \leq \liminf [E(Y_n)] = \liminf [E(X_n)] \quad \square$$